

# Large Deviations

We first introduce some general definitions. Suppose  $S$  is a complete, separable metric space and  $\nu_n$  is a sequence of probability measures on  $(S, \mathcal{S})$ .

Recall that a function  $f : S \mapsto [-\infty, \infty]$  is a lower semicontinuous function if

$$f(x) \leq \liminf_{y \rightarrow x} f(y), \quad \forall x \in S.$$

$f$  is lower semicontinuous if and only if for any real number  $c$ ,  $\{x : f(x) > c\}$  is an open subset of  $S$ . From this one can easily show that if  $\{f_a(x) : a \in A\}$  is a family of continuous functions on  $S$ , then the function

$$f(x) = \sup_{a \in A} f_a(x), \quad x \in S$$

is a lower semicontinuous function on  $S$ .

**Definition 1** A rate function  $I$  is a lower semicontinuous function  $I : S \mapsto [0, \infty]$ . A good rate function is a rate function  $I$  for which all level sets  $\Psi_I(\alpha) = \{x : I(x) \leq \alpha\}$ ,  $\alpha \geq 0$ , are compact subsets of  $S$ .

**Definition 2** We say that  $\{\nu_n\}$  satisfies the large deviation principle with a rate function  $I$  if

(a) for any closed subset  $F$  of  $S$ ,

$$\limsup \frac{1}{n} \log \nu_n(F) \leq - \inf_{x \in F} I(x);$$

(b) for any open subset  $G$  of  $S$ ,

$$\liminf \frac{1}{n} \log \nu_n(G) \geq - \inf_{x \in G} I(x).$$

Suppose that  $\{\nu_n\}$  satisfies the large deviation principle with a rate function  $I$ . If  $\Gamma \in \mathcal{S}$  is a continuity set of  $I$ , i.e.,

$$\inf_{x \in \Gamma^o} I(x) = \inf_{x \in \Gamma} I(x) := I_\Gamma,$$

then

$$\lim_{n \rightarrow \infty} \frac{1}{n} \log \nu_n(\Gamma) = -I_\Gamma$$

which can be written heuristically as

$$\nu_n(\Gamma) \asymp \exp(-nI_\Gamma).$$

Now suppose that  $X_1, X_2, \dots$  are iid random variables. Let  $\mu = EX_1$  and let  $\nu$  be the distribution of  $X_1$ . Let  $\nu_n$  be the distribution of  $\frac{S_n}{n} = \frac{X_1 + \dots + X_n}{n}$ . The logarithmic moment generating function associated with  $\nu$  is

$$\Lambda(\lambda) := \log M(\lambda) := \log E(e^{\lambda X_1}), \quad \lambda \in \mathbf{R}.$$

Obviously,  $\Lambda(0) = 0$ ,  $\Lambda(\lambda) > -\infty$  for any  $\lambda$ . It is possible to have  $\Lambda(\lambda) = \infty$ .

**Definition 3** *The Legendre-Fenchel transform of  $\Lambda(\lambda)$  is*

$$\Lambda^*(x) = \sup_{\lambda \in \mathbf{R}} \{\lambda x - \Lambda(\lambda)\}.$$

From the definition one can easily see that  $\Lambda^*$  is a nonnegative lower semicontinuous function on  $R$ , so it is a rate function.

**Theorem 1** *(Cramer's Theorem) The sequence of probability measures  $\{\nu_n\}$  satisfies the large deviation principle with the convex rate function  $\Lambda^*(\cdot)$ , namely*

(a) *for any closed subset  $F$  of  $\mathbf{R}$ ,*

$$\limsup \frac{1}{n} \log \nu_n(F) \leq - \inf_{x \in F} I(x); \tag{1}$$

(b) *for any open subset  $G$  of  $\mathbf{R}$ ,*

$$\liminf \frac{1}{n} \log \nu_n(G) \geq - \inf_{x \in G} I(x). \tag{2}$$

First we state the properties of  $\Lambda^*(\cdot)$  and  $\Lambda$  that are needed for proving Cramer's Theorem.

**Lemma 2** (a)  *$\Lambda$  and  $\Lambda^*$  are both convex functions on  $R$ .*

(b) *If  $\mathcal{D}_\Lambda := \{\lambda : \Lambda(\lambda) < \infty\} = \{0\}$ , then  $\Lambda^*$  is identically zero. If  $\Lambda(\lambda) < \infty$  for some  $\lambda > 0$ , then  $\mu < \infty$  and for all  $x \geq \mu$ ,*

$$\Lambda^*(x) = \sup_{\lambda \geq 0} \{\lambda x - \Lambda(\lambda)\}, \tag{3}$$

is a nondecreasing function on  $(\mu, \infty)$ . Similarly, if  $\Lambda(\lambda) < \infty$  for some  $\lambda < 0$ , then  $\mu > -\infty$  and for all  $x \leq \mu$ ,

$$\Lambda^*(x) = \sup_{\lambda \leq 0} \{\lambda x - \Lambda(\lambda)\}, \quad (4)$$

is a nonincreasing function on  $(-\infty, \mu)$ . When  $\mu$  is finite,  $\Lambda^*(\mu) = 0$ , and we always have

$$\inf_{x \in \mathbf{R}} \Lambda^*(x) = 0. \quad (5)$$

(c)  $\Lambda(\cdot)$  is differentiable in  $\mathcal{D}_\Lambda^\circ$  with

$$\Lambda'(\eta) = \frac{1}{M(\eta)} E(X_1 e^{\eta X_1}) \quad (6)$$

and

$$\Lambda'(\eta) = y \implies \Lambda^*(y) = \eta y - \Lambda(\eta). \quad (7)$$

**Proof.** (a) The convexity of  $\Lambda$  follows from Hölder's inequality

$$\begin{aligned} \Lambda(\theta\lambda_1 + (1-\theta)\lambda_2) &= \log E[(e^{\lambda_1 X_1})^\theta (e^{\lambda_2 X_1})^{1-\theta}] \\ &\leq \log \{E(e^{\lambda_1 X_1})^\theta E(e^{\lambda_2 X_1})^{1-\theta}\} \\ &= \theta\Lambda(\lambda_1) + (1-\theta)\Lambda(\lambda_2). \end{aligned}$$

The convexity of  $\Lambda^*$  follows from its definition

$$\begin{aligned} \theta\Lambda^*(x_1) + (1-\theta)\Lambda^*(x_2) &= \sup_{\lambda \in \mathbf{R}} \{\theta\lambda x_1 - \theta\Lambda(\lambda)\} + \sup_{\lambda \in \mathbf{R}} \{(1-\theta)\lambda x_2 - (1-\theta)\Lambda(\lambda)\} \\ &\geq \sup_{\lambda \in \mathbf{R}} \{(\theta x_1 + (1-\theta)x_2)\lambda - \Lambda(\lambda)\} \\ &= \Lambda^*(\theta x_1 + (1-\theta)x_2). \end{aligned}$$

(b) If  $\mathcal{D}_\Lambda = \{0\}$ , then  $\Lambda^*(x) = \Lambda(0) = 0$  for all  $x \in \mathbf{R}$ . If  $\Lambda(\lambda) = \log M(\lambda) < \infty$  for some  $\lambda > 0$ , then  $\int_0^\infty x d\nu \leq M(\lambda)/\lambda < \infty$ , implying  $\mu = \int x d\nu < \infty$ . Now, for all  $\lambda \in \mathbf{R}$ , by Jensen's inequality

$$\Lambda(\lambda) = \log E(e^{\lambda X_1}) \geq E(\log e^{\lambda X_1}) = \lambda\mu.$$

If  $\mu = -\infty$ , then  $\Lambda(\lambda) = \infty$  for  $\lambda$  negative and (3) holds trivially. When  $\mu$  is finite, it follows from the preceding inequality that  $\Lambda^*(\mu) = 0$ . In this case, for any  $x \geq \mu$  and every  $\lambda < 0$ ,

$$\lambda x - \Lambda(\lambda) \leq \lambda\mu - \Lambda(\lambda) \leq \Lambda^*(\mu) = 0,$$

and (3) follows. Observe that (3) implies the monotonicity of  $\Lambda^*$  on  $(\mu, \infty)$ , since for any  $\lambda \geq 0$ ,  $\lambda x - \Lambda(\lambda)$  is a nondecreasing function of  $x$ .

When  $\Lambda(\lambda) < \infty$  for some  $\lambda < 0$ , both (4) and the monotonicity of  $\Lambda^*$  on  $(-\infty, \mu)$  follows by considering the logarithmic moment generating function of  $-X_1$ , for which the preceding proof applies.

It remains to prove that  $\inf_{x \in \mathbf{R}} \Lambda^*(x) = 0$ . This is already established for  $\mathcal{D}_\Lambda = \{0\}$ , in which case  $\Lambda^* \equiv 0$ , and when  $\mu$  is finite, in which case,  $\Lambda^*(\mu) = 0$ . Now consider the case  $\mu = -\infty$  while  $\Lambda(\lambda) < \infty$  for some  $\lambda > 0$ . Then by (3) and Chebyshev

$$\begin{aligned} \log \nu([x, \infty)) &\leq \inf_{\lambda \geq 0} \log E(e^{\lambda(X_1 - x)}) \\ &= -\sup_{\lambda \geq 0} \{\lambda x - \Lambda(\lambda)\} = -\Lambda^*(x). \end{aligned}$$

Hence

$$\lim_{x \rightarrow -\infty} \Lambda^*(x) \leq \lim_{x \rightarrow -\infty} \{-\log \nu([x, \infty))\} = 0,$$

and (5) follows.

The last remaining case, that of  $\mu = \infty$  while  $\Lambda(\lambda) < \infty$  for some  $\lambda < 0$ , is settled by considering the logarithmic moment generating function of  $-X_1$ .

(c) The identity (6) follows by interchanging the order of differentiation and integration. This is justified by the dominated convergence theorem since  $f_\epsilon(x) := (e^{(\eta+\epsilon)x} - e^{\eta x})/\epsilon$  converges pointwise to  $x e^{\eta x}$  as  $\epsilon \rightarrow 0$ , and  $|f_\epsilon(x)| \leq e^{\eta x}(e^{\delta|x|} - 1)/\delta := h(x)$  for every  $\epsilon \in (-\delta, \delta)$ , while  $E|h(X_1)| < \infty$  for  $\delta$  small enough.

Let  $\Lambda'(\eta) = y$  and consider the function  $g(\lambda) := \lambda y - \Lambda(\lambda)$ . Since  $g(\cdot)$  is a concave function and  $g'(\eta) = 0$ , it follows that  $g(\eta) = \sup_{\lambda \in \mathbf{R}} g(\lambda)$  and (7) is established.

**Proof of Cramer's Theorem** (a) Let  $F$  be a nonempty closed subset of  $\mathbf{R}$ . (1) holds trivially when  $I_F := \inf_{x \in F} \Lambda^*(x) = 0$ . Assume  $I_F > 0$ . It follows from part (b) of the lemma that  $\mu$  exists. For all  $x$  and every  $\lambda > 0$ , Chebyshev yields

$$\begin{aligned} \nu_n([x, \infty)) &= E(1_{\{\frac{S_n}{n} - x \geq 0\}}) \\ &\leq E(e^{n\lambda(\frac{S_n}{n} - x)}) \\ &= e^{-n\lambda x} \prod_{i=1}^n E(e^{\lambda X_i}) \\ &= e^{-n(\lambda x - \Lambda(\lambda))} \end{aligned} \tag{8}$$

Therefore, if  $\mu < \infty$ , then by (3) for every  $x > \mu$ ,

$$\nu_n([x, \infty)) \leq e^{-n\Lambda^*(x)}. \tag{9}$$

By a similar argument, if  $\mu > -\infty$  and  $x < \mu$ , then

$$\nu_n((-\infty, x]) \leq e^{-n\Lambda^*(x)}. \tag{10}$$

First, consider the case of  $\mu$  finite. Then  $\Lambda^*(\mu) = 0$ , and because by assumption  $I_F > 0$ ,  $\mu$  must be in  $F^c$ . Let  $(x_-, x_+)$  be the component of  $F^c$  containing  $\mu$ . Note that  $x_- < x_+$  and either  $x_-$  or  $x_+$  must be finite since  $F$  is nonempty. If  $x_-$  is finite, then  $x_- \in F$ , and

consequently  $\Lambda^*(x_-) \geq I_F$ . Likewise,  $\Lambda^*(x_+) \geq I_F$  whenever  $x_+$  is finite. Applying (9) for  $x = x_+$  and (10) for  $x = x_-$  we get

$$\nu_n(F) \leq \nu_n(-\infty, x_-] + \nu_n([x_+, \infty)) \leq 2e^{-nI_F}$$

and the upper bounds follows.

Suppose now that  $\mu = -\infty$ . Then, since  $\Lambda^*$  is nondecreasing, it follows from (5) that  $\lim_{x \rightarrow -\infty} \Lambda^*(x) = 0$ , and hence  $x_+ = \inf\{x : x \in F\}$  is finite for otherwise  $I_F = 0$ . Since  $F$  is closed,  $x_+ \in F$  and consequently  $\Lambda^*(x_+) \geq I_F$ . Moreover,  $F \subset [x_+, \infty)$  and, therefore, the large deviation upper bounds follows by applying (9) for  $x = x_+$ .

The case of  $\mu = \infty$  is handled analogously.

(b) We prove next that for every  $\delta > 0$

$$\liminf \frac{1}{n} \log \nu_n((-\delta, \delta)) \geq \inf_{\lambda \in \mathbf{R}} \Lambda(\lambda) = -\Lambda^*(0). \quad (11)$$

Since the transform  $Y = X - x$  results in  $\Lambda_Y(\lambda) = \Lambda(\lambda) - \lambda x$  and hence  $\Lambda_Y^*(\cdot) = \Lambda^*(\cdot + x)$ , it follows from the proceeding inequality that for every  $x$  and every  $\delta > 0$ ,

$$\liminf \frac{1}{n} \log \nu_n((x - \delta, x + \delta)) \geq -\Lambda^*(x). \quad (12)$$

For any open subset  $G$  of  $\mathbf{R}$  and any  $x \in G$ , and any  $\delta > 0$  small enough,  $(x - \delta, x + \delta) \subset G$ . Thus the large deviation lower bounds follows from (12).

Turning to the proof of the key inequality (11), first suppose that  $\nu((-\infty, 0)) > 0$ ,  $\nu((0, \infty)) > 0$  and  $\nu$  is supported by a bounded subset of  $\mathbf{R}$ . By the former assumption,  $\Lambda(\lambda) \rightarrow \infty$  as  $|\lambda| \rightarrow \infty$ , and by the later assumption,  $\Lambda(\cdot)$  is finite everywhere. Accordingly,  $\Lambda(\cdot)$  is a continuous, differentiable function, and hence there exists a finite  $\eta$  such that  $\Lambda(\eta) = \inf_{\lambda \in \mathbf{R}} \Lambda(\lambda)$  and  $\Lambda'(\eta) = 0$ . Define a new probability measure  $\tilde{\nu}$  in terms of  $\nu$  by

$$\frac{d\tilde{\nu}}{d\nu}(x) = e^{\eta x - \Lambda(\eta)}$$

and observe that  $\tilde{\nu}$  is a probability measure because

$$\int_{\mathbf{R}} d\tilde{\nu} = \frac{1}{M(\eta)} \int e^{\eta x} d\nu = 1.$$

Let  $\tilde{\nu}_n$  be the distribution of  $\frac{S_n}{n}$  when  $X_1, X_2, \dots$  are iid with distribution  $\tilde{\nu}$ . Note that for every  $\epsilon > 0$ ,

$$\begin{aligned} \nu_n((-\epsilon, \epsilon)) &= \int_{|\sum_{i=1}^n x_i| < n\epsilon} \nu(dx_1) \cdots \nu(dx_n) \\ &\geq e^{-n\epsilon|\eta|} \int_{|\sum_{i=1}^n x_i| < n\epsilon} \exp\left(\eta \sum_{i=1}^n x_i\right) \nu(dx_1) \cdots \nu(dx_n) \\ &= e^{-n\epsilon|\eta|} e^{n\Lambda(\eta)} \tilde{\nu}_n((-\epsilon, \epsilon)) \end{aligned} \quad (13)$$

By (6) and the choice of  $\eta$

$$E_{\tilde{\nu}}(X_1) = \frac{1}{M(\eta)} \int x e^{\eta x} d\nu = \Lambda'(\eta) = 0.$$

Hence, by the weak law of large numbers,

$$\lim_{n \rightarrow \infty} \tilde{\nu}_n((-\epsilon, \epsilon)) = 1. \quad (14)$$

It now follows from (13) that for every  $0 < \epsilon < \delta$ ,

$$\begin{aligned} \liminf \frac{1}{n} \log \nu_n((-\delta, \delta)) &\geq \liminf \frac{1}{n} \log \nu_n((-\epsilon, \epsilon)) \\ &\geq \Lambda(\eta) - \epsilon|\eta| \end{aligned}$$

and (11) follows by letting  $\epsilon \rightarrow 0$ .

Suppose that  $\nu$  is of unbounded support, while both  $\nu((-\infty, 0)) > 0$  and  $\nu((0, \infty)) > 0$ . Fix  $M$  large enough so that  $\nu([-M, 0)) > 0$  and  $\nu((0, M]) > 0$ , and let

$$\Lambda_M(\lambda) = \log \int_{-M}^M e^{\lambda x} d\nu.$$

Let  $\bar{\nu}$  be the probability measure defined by

$$\bar{\nu}(A) = \frac{\nu(A \cap [-M, M])}{\nu([-M, M])},$$

and let  $\bar{\nu}_n$  be the distribution of  $\frac{S_n}{n}$  when  $X_1, X_2, \dots$  are iid with distribution  $\bar{\nu}$ . The for all  $n$  and every  $\delta > 0$ ,

$$\nu_n((-\delta, \delta)) \geq \bar{\nu}_n((-\delta, \delta)) \nu_n([-M, M])^n.$$

Observe that by the proceeding paragraph, (11) holds for  $\bar{\nu}_n$ . Therefore, with the logarithmic moment generating function associated with  $\bar{\nu}$  being  $\Lambda_M(\lambda) - \log \nu([-M, M])$ ,

$$\begin{aligned} \liminf \frac{1}{n} \log \nu_n((-\delta, \delta)) &\geq \log \nu([-M, M]) + \liminf \frac{1}{n} \log \bar{\nu}_n((-\delta, \delta)) \\ &\geq \inf_{\lambda \in \mathbf{R}} \Lambda_M(\lambda). \end{aligned}$$

With  $I_M = -\inf_{\lambda \in \mathbf{R}} \Lambda_M(\lambda)$  and  $I^* = \limsup_{M \rightarrow \infty} I_M$ , it follows that

$$\liminf \frac{1}{n} \log \nu_n((-\delta, \delta)) \geq -I^*. \quad (15)$$

Note that  $\Lambda_M(\cdot)$  is nondecreasing in  $M$ , and so is  $-I_M$ . Moreover,  $-I_M \leq \Lambda_M(0) \leq \Lambda(0) = 0$ , and hence  $-I^* \leq 0$ . Now, since  $-I_M$  is finite for all  $M$  large enough,  $-I^* > -\infty$ . Therefore the level sets  $\{\lambda : \Lambda_M(\lambda) \leq -I^*\}$  are nonempty, compact sets that are decreasing with respect to  $M$ , and hence there exists at least one point, denoted  $\lambda_0$ , in their intersection.

By the monotone convergence theorem,  $\Lambda(\lambda_0) = \lim \Lambda_M(\lambda_0) \leq -I^*$ , and consequently the bounds (15) yields (11), now for  $\nu$  of unbounded support.

The proof of (11) for an arbitrary  $\mu$  is completed by observing that if either  $\mu((-\infty, 0)) = 0$  or  $\mu((0, \infty)) = 0$ , then  $\Lambda$  is a monotone function with  $\inf_{\lambda \in \mathbf{R}} \Lambda(\lambda) = \log \nu(\{0\})$ . Hence, in this case, (11) follows from

$$\nu_n((-\delta, \delta)) \geq \nu_n(\{0\}) = \nu(\{0\})^n.$$

**Lemma 3** *If  $0 \in \mathcal{D}_\Lambda^o$  then  $\Lambda^*$  is a good rate function. Moreover, if  $\mathcal{D}_\Lambda = \mathbf{R}$ , then*

$$\lim_{|x| \rightarrow \infty} \frac{\Lambda^*(x)}{|x|} = \infty. \quad (16)$$

**Proof.** As  $0 \in \mathcal{D}_\Lambda^o$ , there exist  $\lambda_- < 0$  and  $\lambda_+ > 0$  that are both in  $\mathcal{D}_\Lambda$ . Since for any  $\lambda \in \mathbf{R}$ ,

$$\frac{\Lambda^*(x)}{|x|} \geq \lambda \text{sign}(x) - \frac{\Lambda(\lambda)}{|x|},$$

it follows that

$$\liminf_{|x| \rightarrow \infty} \frac{\Lambda^*(x)}{|x|} \geq \min\{\lambda_+, -\lambda_-\} > 0.$$

In particular,  $\Lambda^*(x) \rightarrow \infty$  as  $|x| \rightarrow \infty$ , and its level sets are closed and bounded, hence compact. Thus  $\Lambda^*$  is a good rate function. Note that (16) follows for  $\mathcal{D}_\Lambda = \mathbf{R}$  by considering  $-\lambda_- = \lambda_+ \rightarrow \infty$ .

### Examples

- (a) If the iid sequence  $X_1, X_2, \dots$  has the normal distribution with mean  $\mu$  and variance  $\sigma^2$  as their common distribution, then  $M(\lambda) = \exp(\mu\lambda + \frac{\sigma^2\lambda^2}{2})$ ,  $\Lambda(\lambda) = \mu\lambda + \frac{\sigma^2\lambda^2}{2}$ , and  $\Lambda^*(x) = \frac{(x-\mu)^2}{2\sigma^2}$ .
- (b) If the iid sequence  $X_1, X_2, \dots$  has the exponential distribution with parameter  $\theta > 0$  as their common distribution, then  $M(\lambda) = \frac{\theta}{\theta-\lambda}$  for  $\lambda < \theta$ , and  $\infty$  otherwise;  $\Lambda(\lambda) = \log \frac{\theta}{\theta-\lambda}$  for  $\lambda < \theta$ , and  $\infty$  otherwise;  $\Lambda^*(x) = \theta x - 1 - \log(\theta x)$  for  $x > 0$ , and  $\infty$  otherwise.
- (c) If the iid sequence  $X_1, X_2, \dots$  has the Bernoulli distribution with parameter  $p \in (0, 1)$  as their common distribution, then  $M(\lambda) = pe^\lambda + 1 - p$ ,  $\Lambda^*(x) = x \log(\frac{x}{p}) + (1-p) \log(\frac{1-x}{1-p})$  for  $x \in [0, 1]$  and  $\infty$  otherwise.
- (d) If the iid sequence  $X_1, X_2, \dots$  has the Poisson distribution with parameter  $\theta > 0$  as their common distribution, then  $M(\lambda) = \exp(\theta(e^\lambda - 1))$  and  $\Lambda^*(x) = \theta - x + x \log(\frac{x}{\theta})$  for  $x \geq 0$  and  $\infty$  otherwise.